

## Optimization of *Dothideomyces sp.* NRC-SSW Chitosanase Productivity and Activity Using Response Surface Methodology

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**A**MONG the methods used for the hydrolysis of chitosan, enzymatic hydrolysis using chitosanase was selected in order to produce high yield of specific chitooligosaccharides with less environmental pollutions. The production of *Dothideomyces sp.* NRC-SSW extracellular chitosanase was statistically optimized in which a two - phase experimental design was applied. Plackett – Burman design was used to evaluate the relative importance of culture conditions and medium components for chitosanase production. Chitosan concentration, agitation speed and incubation period were found to be the most significant variables that affected the chitosanase production and their optimal values were obtained by applying Box-Behnken design. The optimized medium composed of (g/L) chitosan, 30; K<sub>2</sub>HPO<sub>4</sub>, 1.5; MgSO<sub>4</sub>, 0.4; KCl, 4.0; yeast extract, 18.5 and FeSO<sub>4</sub>, 0.01; at pH 5.5, 30°C and 180rpm for 96h gave 13.9U/mL with 36.3% increase in the activity. The R<sup>2</sup> value was 0.954 and this indicated the aptness of the model. The optimization of the hydrolytic conditions required for chitooligosaccharides production was also performed by Box-Behnken design. The highest yield of chitooligosaccharides was obtained with enzyme/ substrate ratio 0.05U/mg in 0.2M Tris HCl buffer incubation at 60°C for 5h. The cytotoxic activity of the chitooligosaccharides was tested in vitro against Hep-G2 and MCF7.

**Keywords:** *Dothideomyces sp.*, Chitosanase, Plackett-Burman design, Box-Behnken design, Chitooligosaccharides, Anticancer.

### Introduction

Chitooligosaccharides (COS) are water soluble derivatives of chitosan that composed mainly of  $\beta$ -(1-4)-linked D-glucosamine with randomly distributed N-acetyl-D-glucosamine. They have been reported to have remarkable biological activities as well as their potential applications in food, pharmaceutical and agricultural industries [1, 2]. Recently COS gained a great attention due to their various health benefits as antitumor, anti-inflammatory, anti-hypertensive and hypocholesterimic effects [3, 4]. These various applications attract the research focus to improve the COS productivity to meet the industrial requirements with minimal environmental pollution.

Enzymatic hydrolysis of chitosan to COS has several advantages as it gave specific and high yield of COS with less environmental pollution

[5-7]. Chitosanases are the enzymes that used mainly for hydrolysis of chitosan to COS but their applications are limited due to the high cost and low levels of their production [4]. Several interests are focused to optimize the chitosanase productivity.

Response Surface methodology (RSM) is a statistical technique used successfully in the optimization of complex chemical, biochemical and food process. This technique has received much attention in the investigation of the optimization process for the production of microbial enzymes [7-10].

Fungal chitosanases are poorly studied in comparison to bacterial chitosanases [11, 12]. So in this research a fungal strain identified using 18S rRNA was used for chitosanase production. Statistical optimization with a two level experimental design was carried out as

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follow, Plackett – Burman design was applied to identify the most significant factors that influence the chitosanase production then followed by Box-Behnken design to optimize the variables of the highest effect. The optimal conditions required for the production of COS through the hydrolysis of chitosan by the produced enzyme was investigated by applying Box-Behnken design. Also the anticancer activity of the produced COS was evaluated.

## **Experimental**

### *Materials*

Chitosan (low and medium molecular weight), glucosamine hydrochloride and yeast extract were obtained from Sigma-Aldrich, Saint Louis, USA. Dinitrosalicylic acid (DNS) was obtained from Panreac, Barcelona, Spain. TLC Silica gel 60 plate was obtained from Merck, Darmstadt, Germany. All other chemicals were of analytical or HPLC grade.

### *Microorganism*

The fungal strain used in this research was isolated from shrimp shell wastes collected from the local market. The morphological features of the isolated strain were studied under light microscope and scanning electron microscope. Molecular identification of the strain was done by Sigma Scientific Services Co. as follow: DNA extraction was done according to the protocol of GeneJet Plant genomic DNA purification using Thermo kit (K0791). PCR was made using Maxima Hot Start PCR Master Mix (thermo K0221) using ITS1 and ITS4 primers. Thermo-cycling process was done with initial denaturation at 95°C for 10min, denaturation at 95°C for 30sec, annealing at 55°C for 1min, extension at 72°C for 1min and then final extension at 72°C for 15min for 35 cycles. The PCR product was purified using GeneJET™ PCR purification thermo kit (K0701). Sequencing of the PCR product was done in GATC Company using ABI 3730xl DNA sequencer using forward and reverse primers.

### *Chitosanase production*

#### *Preparation of soluble chitosan for culture medium*

Ten gram of powder chitosan (medium molecular weight) was suspended and dissolved through stirring in 1L of 0.1M HCl solution (pH 1.5) then the pH of the solution was adjusted to 5.5 using NaOH (10N) [13].

#### *Inoculum preparation*

An inoculum culture was obtained by culturing

the spore suspension of the fungal strain in 250mL Erlenmeyer flask contain 90mL of sterilized medium composed of (g/L) chitosan (prepared as indicated above) 2.0; K<sub>2</sub>HPO<sub>4</sub> 1.0; MgSO<sub>4</sub> 0.5; KCl 5.0; yeast extract 3.0; peptone 5.0; NaNO<sub>3</sub> 2.0 and FeSO<sub>4</sub> 0.01 at initial pH 5 incubated at 30°C and 180rpm for 48h [6]. The resulted pre-inoculated medium was further used to culture the fermentation medium.

### *Fermentation process*

The fermentation was performed in 250mL Erlenmeyer flask contained 90mL of the basal fermentation medium resulted after single factor optimization [6] that consisted of (g/L): soluble chitosan, 20; K<sub>2</sub>HPO<sub>4</sub>, 1.0; MgSO<sub>4</sub>, 0.3; KCl, 4.0; yeast extract, 15.0; FeSO<sub>4</sub>, 0.01; at pH 5 before autoclaving. The flasks were cultured by 10ml of the pre-inoculated medium and incubated at 30°C in an incubator at 180rpm for 3days [6]. At the end of the culture period, the fermented medium was centrifuged at 5000rpm for 10min. The clear culture filtrate was used to estimate the chitosanase activity.

### *Enzyme assay*

Chitosanase activity was assayed using low molecular weight chitosan as a substrate. Chitosan (1%) was treated with 1M acetic acid and adjusted to pH 5 using 2M sodium acetate before its use as a substrate for enzyme assay [14]. The enzyme activity was determined according to Pagnoncelli *et al.*, [15] by mixing 500µL of the clear culture filtrate with 500µL of 1% soluble chitosan and the reaction mixture was incubated at 50°C for 30min. At the end of the assay time, 2.5mL of DNS was added to stop the reaction. The above reaction mixture with heat inactivated enzyme was used as a blank. The reducing sugars were measured immediately by using DNS method [16] with D-glucosamine as the standard. One unit of chitosanase was defined as the amount of enzyme that released 1µmol of D-glucosamine per minute under the assay conditions.

### *Statistical optimization of chitosanase production*

The optimization study was performed by Plackett – Burman design followed by Box-Behnken design.

### *Plackett – Burman design*

For multivariable processes such as biochemical systems, in which numerous potentially influential factors were involved, it is necessary to analyze the process with an initial screening design prior to optimization

using Box and Behnken [17]. Plackett–Burman experimental design [18] was used to evaluate the relative importance of culture conditions and medium components for the production of chitosanase in submerged fermentation (SmF). Seven independent variables were screened in eight trials organized according to the plackett – Burman design matrix described in the results section. For each variable, a high (+) and low (-) level was tested.

Plackett–Burman experimental design was based on the first order linear model:

$$Y = B_0 + \sum B_i X_i \quad \text{Eq. (1)}$$

Where Y is the response (chitosanase production),  $B_0$  is the model intercept and  $B_i$  is the variables estimates. The main effect of each variable was determined by the following equation:

$$E_{(X_i)} = \frac{(\sum M_i^+ - M_i^-)}{N} \quad \text{Eq. (2)}$$

Where  $E_{(X_i)}$  is the effect of the tested variable.  $M_i^+$  and  $M_i^-$  represent chitosanase production from the trials where the independent variable (Xi) measured was present at high and low concentrations, respectively and N is the number of trials.

The standard error (SE) of the concentration effect was the square root of the variance of an effect, and the significance level (p-value) of each concentration effect was determined using student's t-test

$$t(X_i) = E(X_i)/SE \quad \text{Eq. (3)}$$

Where  $E(X_i)$  is the effect of variable Xi.

#### Box-Behnken Design

In order to describe the nature of response surface in the experimental region, Box-Behnken design [17] was applied. In this model, the most significant independent variables extracted from the above experiment namely  $X_1$ ,  $X_2$  and  $X_3$  are included and each factor examined at three different levels, low (-), high (+) and control or basal (0). For predicting the optimal point, a second-order polynomial function was fitted to correlate relationship between independent variables and response (chitosanase biosynthesis). For the three factors the polynomial equation is in the following form:

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_3 + \beta_{11} X_1^2 + \beta_{22} X_2^2 + \beta_{33} X_3^2 + \beta_{12} X_1 X_2 + \beta_{13} X_1 X_3 + \beta_{23} X_2 X_3 \quad \text{Eq. (4)}$$

Where Y is the predicted response(chitosanase

production) and  $X_1$ ,  $X_2$  and  $X_3$  are the independent variables corresponding to the concentration of carbon source, shaking speed (rpm) and incubation period respectively;  $\beta_0$  is the intercept regression coefficient,  $\beta_1$ ,  $\beta_2$ ,  $\beta_3$  are linear coefficients,  $\beta_{11}$ ,  $\beta_{22}$ ,  $\beta_{33}$  are quadratic coefficients,  $\beta_{12}$ ,  $\beta_{13}$ ,  $\beta_{23}$  are cross product coefficients.

#### Partial purification of chitosanase

This was done by subjecting the crude enzyme (culture filtrate) to fractional precipitation with ethanol at 30-90% ethanol concentration with 10% intervals. Each fraction was assayed for chitosanase activity and protein content [19].

#### Determination of protein content

The protein content of the enzyme preparations was determined by the method of Lowry *et al.*, [20] using bovine serum albumin as standard.

#### Hydrolysis of chitosan

Soluble chitosan (2%w/v) was prepared by dissolving the powder (medium molecular weight) in 1M acetic acid then adjusted to pH 4.5 (using 0.2M Tris HCl buffer). The reaction mixture contained 2mL of 2% soluble chitosan and different enzyme concentration to reach to enzyme/ substrate ratio (E/S) ranged from 0.0125 to 0.2U/mg. The reaction was carried out in a shaking water bath at 55°C for 4h. The reaction was stopped by keeping the mixture for 10min in boiling water bath. The hydrolysis products of these reactions were identified by silica gel thin-layer chromatography plate using a mixture of propanol: water: ammonia (7: 2: 1 v/v) as a mobile phase [21]. The amino sugars were visualized with diphenyl amine-aniline reagent [22].

#### Statistical optimization of the hydrolytic conditions

Box-Behnken design was applied to verify the most suitable conditions to hydrolyze the chitosan to COS. Box-Behnken design was applied in 27-run with four factors (Reaction time, temperature, enzyme units and substrate concentration (%)) and three levels (low (-), high (+) and control or basal (0) including three replicates at the center point and the results were expressed as the amount of reducing sugars released (%) as reported by Sun *et al.*, [23]. The second-order polynomial equation used to calculate the predicted response was as follow:

$$Y = B_0 + \sum B_i X_i + \sum B_{ii} X_i^2 + \sum B_{ij} \sum X_i X_j \quad \text{Eq. (5)}$$

Where Y represents response variable,  $\beta_0$  is the

interception coefficient,  $\beta_i$  the coefficient of the linear effect,  $\beta_{ii}$  the coefficient of quadratic effect and  $\beta_{ij}$  are cross product coefficients,  $X_i, X_j$  are independent variables which influence the response variable Y.

#### Statistical analysis

All experiments were performed in triplicates and the averages of the results were reported. The results were statistically analyzed by using the analysis of Variance (ANOVA) techniques, in which p value of  $\leq 0.05$  was regarded as significant. Statistical software SPSS (version 16.0) was used for the regression analysis of the experimental data obtained. The quadratic models were represented as contour plots (3D) and response surface curves were generated by using STATISTICA (0.6).

#### Anticancer activity of the produced COS

Cytotoxic activities of the produced COS were tested on Human liver cancer cell line (HEP-G2) and Breast cancer cell line (MCF7) according to the method of Skehan *et al.* [24] and the results were expressed as  $IC_{50}$  ( $\mu\text{g/ml}$ ) which represented the concentration of the pure chitoooligosaccharide (eluted from TLC plates) that reduced survival to 50%. In this method the cells were plated in 96 multi well plate ( $10^4$  cells/well) for 24h to allow attachment of the cells to the wall of the plate then different concentrations of the COS (0, 12.5, 25.50 and  $100\mu\text{g/mL}$ ) were added to the cell monolayer in which triplicate wells were prepared for each individual dose. After incubation in atmosphere of 5%  $\text{CO}_2$  for 48h the cells were fixed, washed and stained with Sulfo-Rhodamine-B stain. The color intensity was measured in an ELISA reader.

## Results

#### Identification of the chitosanase producing fungal strain

The growth of the fungus on potato dextrose agar (PDA) showed that the color of the fungal spores was dull green. The morphological features of the strain using light microscope and scanning electron microscope were shown in Fig. 1. The partial sequence of the 18S DNA showed 99% similarity with *Dothideomycetes* sp. The phylogenetic tree (Fig. 2) also showed that the strain was closely related to *Dothideomycetes* sp. So the fungus used in this research was identified as *Dothideomycetes* sp. NRC-SSW.

#### Optimization of chitosanase production by RSM

The optimization of the medium composition and the culture conditions for extracellular chitosanase

production was carried out by statistical methodology.

#### Plackett – Burman design

The mean response as the chitosanase activity obtained with the observed response generated in Plackett – Burman was presented in Table 1. The data indicated that there was a wide variation in extracellular chitosanase activity ranged from 1.86 to 11.0U/mL that reflected the importance of medium optimization to attain the maximum chitosanase production. The analysis of the data from Plackett – Burman experiment involved a first order (main effects) model. The main effects of the examined factors on the extracellular chitosanase activity were calculated and presented graphically in Fig. 3 offering the view for ranking of factors estimates obtained by Plackett – Burman design. The data in Fig. 3 showed that the extracellular chitosanase production was positively affected by all of the tested factors which indicated that the high concentration of these variables was near to optimum.

The first order model describing the correlation of the seven factors and the chitosanase activity could be presented as follows:

$$Y_{\text{Activity}} = -25.113 + 0.986X_1 + 3.148X_2 + 0.174X_3 + 13.875X_4 + 81.875X_5 + 0.041X_6 + 1.833X_7 \quad \text{Eq. (6)}$$

The t test, p effect and the confidence level were presented in Table 2. The variables (chitosan concentration, agitation speed and incubation period) that showed high confidence level in the Plackett–Burman design were selected for further optimization.

#### Box-Behnken Design

Box-Behnken design was used to reach the optimum concentration of the most significant chosen variables. The coded and un-coded level of the three independent variables investigated at three different level (-, 0, +) with 15 trials listed in Table 3. The table also showed the Box-Behnken design, the observed and the predicted chitosanase activity. Multiple regression analysis of the experimental data (Table 4) gave the following second order polynomial:

$$Y_{\text{Activity}} = -87.938 + 48.125X_1 - 0.056X_2^2 + 22.587X_3^2 - 7.350X_1^2 - 2.737X_3^2 - 0.057X_1X_2 + 0.750X_1X_3 - 0.008X_2X_3 \quad \text{Eq. (7)}$$

Where  $Y_{\text{Activity}}$  was the response (chitosanase production) and  $X_1$ ,  $X_2$  and  $X_3$  were the coded values of the test variables (chitosan concentration, agitation speed and incubation

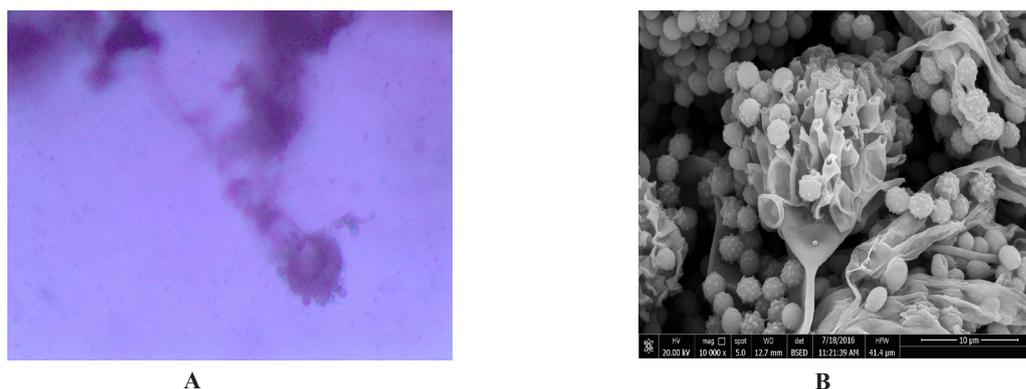


Fig. 1. Morphological features of the fungal strain using light microscope (A) and Scanning electron microscope (B).

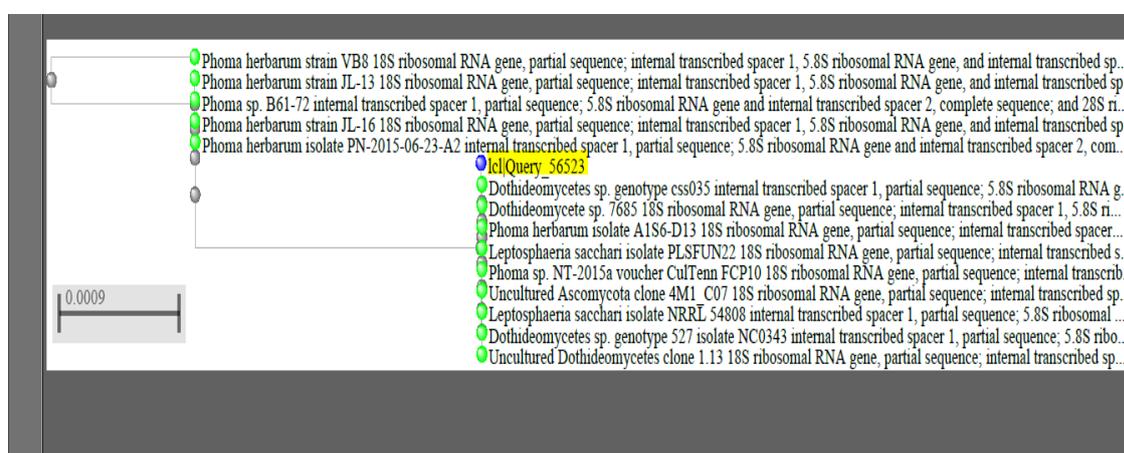


Fig. 2. The phylogenetic tree.

TABLE 1. Plackett – Burman design with coded values and the observed results.

Trial	Incubation time (day)	Chitosan (%)	Yeast extract (%)	$K_2HPO_4$ (%)	$MgSO_4$ (%)	Agitation speed (rpm)	Initial pH	extracellular chitosanase (U/mL)
1	(2) -	(1.5) -	(1.25) -	(0.15) +	(0.04) +	(210) +	(4.5) -	4.09
2	(4) +	(1.5) -	(1.25) -	(0.05) -	(0.02) -	(210) +	(5.5) +	4.87
3	(2) -	(2.5) +	(1.25) -	(0.05) -	(0.04) +	(150) -	(5.5) +	5.20
4	(4) +	(2.5) +	(1.25) -	(0.15) +	(0.02) -	(150) -	(4.5) -	5.09
5	(2) -	(1.5) -	(1.85) +	(0.15) +	(0.02) -	(150) -	(5.5) +	1.86
6	(4) +	(1.5) -	(1.85) +	(0.05) -	(0.04) +	(150) -	(4.5) -	2.25
7	(2) -	(2.5) +	(1.85) +	(0.05) -	(0.02) -	(210) +	(4.5) -	4.27
8	(4) +	(2.5) +	(1.85) +	(0.15) +	(0.04) +	(210) +	(5.5) +	11.00

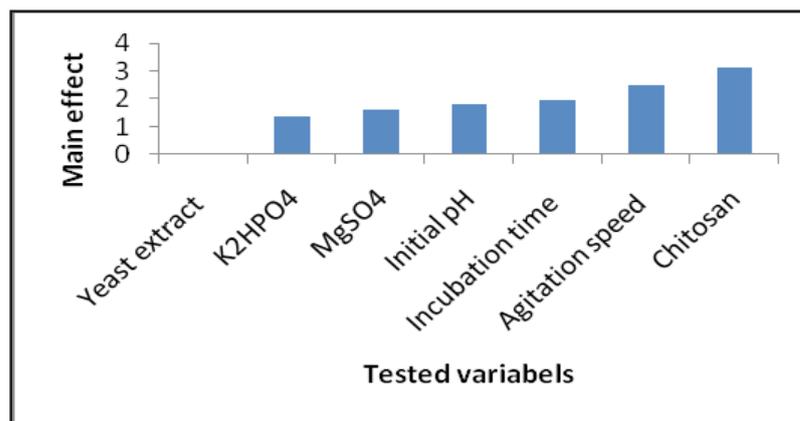


Fig. 3. Main effects of independent variables on extracellular chitosanase production according to the results of the Plackett – Burman experiment.

TABLE 2. Statistical analysis of Plackett- Burman design showing coefficient values, t-and P- values for each variable on extracellular chitosanase production.

Variables	Chitosanase analysis			
	Coefficient	t-statistics	P-value	Confidence level(%)
Intercept	-25.113			
Incubation time (day)	0.986	1.848641	0.0535	94.65
Chitosan(%)	3.148	2.949859	0.0107	98.93
Yeast extract (%)	0.174	0.053889	0.4793	52.07
K <sub>2</sub> HPO <sub>4</sub> (%)	13.875	1.300375	0.1173	88.27
MgSO <sub>4</sub> (%)	81.875	1.534677	0.0844	91.56
Agitation speed	0.041	2.326617	0.0264	97.36
Initial pH	1.833	1.717432	0.0648	93.52

TABLE 3. Examined concentration of the key variables and results of Box-Behnken Design experiments.

Trials	Independent variable			Observed C.f Chitosanase (U/mL)	Predicted C.f Chitosanase (U/mL)
	X <sub>1</sub> Chitosan(%)	X <sub>2</sub> rpm	X <sub>3</sub> Incubation time		
1	2(-)	180(-)	4(0)	8.5	8.957
2	3(+)	180(-)	4(0)	13.9	13.126
3	2(-)	240(+)	4(0)	6.9	7.649
4	3(+)	240(+)	4(0)	8.9	8.4160
5	2(-)	210(0)	3(-)	5.3	4.737
6	3(+)	210(0)	3(-)	5.8	6.455
7	2(-)	210(0)	5(+)	6.3	5.618
8	3(+)	210(0)	5(+)	8.3	8.837
9	2.5(0)	180(-)	3(-)	9.0	9.076
10	2.5(0)	240(+)	3(-)	6.8	6.565
11	2.5(0)	180(-)	5(+)	11.0	11.206
12	2.5(0)	240(+)	5(+)	7.8	7.698
13	2.5(0)	210(0)	4(0)	11.0	10.987
14	2.5(0)	210(0)	4(0)	11.0	10.987
15	2.5(0)	210(0)	4(0)	11.0	10.987

**TABLE 4. Analysis of Box-Behnken Design for extracellular chitosanase production.**

Term	Regression coefficient	Standard error	t- test	P-value
<b>Intercept</b>	-87.938	55.377	-1.588	0.210
$X_1$	48.125	16.048	2.999	0.058
$X_2$	-0.056	.338	-0.165	0.879
$X_3$	22.587	7.066	3.197	0.049
$X_1^2$	-7.350	2.738	-2.685	0.075
$X_2^2$	0.000	0.001	0.566	0.611
$X_3^2$	-2.737	0.684	-3.999	0.028
$X_1X_2$	-0.057	0.034	-1.643	0.199
$X_1X_3$	0.750	1.035	0.725	0.521
$X_2X_3$	-0.008	0.017	-0.483	0.662

F value = 6.966; P>F= 0.0069; R<sup>2</sup>=0.954; R = 0.977; Adjusted R<sup>2</sup>=0.817

period) respectively.

The graphical design of the regression equation was represented by the three-dimensional response surface and the two-dimensional contour plots. They were helpful in understanding both the main and the interaction effects of the factors on the response value. Figure 4a-c showed the response surface and contour plots of chitosan concentration and agitation speed, chitosan concentration and incubation period, also agitation speed and incubation period on chitosanase production respectively, keeping the other variable at the fixed zero level.

The analysis of the results using ANOVA showed a significant F-value (6.966) that indicated the significance of the model. Model terms had values of Prob> F (0.0069) less than 0.05, considered significant. The coefficient (R<sup>2</sup>) for chitosanase activity was 0.954 that indicated the goodness of the model. The closer the R<sup>2</sup> to 1 indicated that the model was useful in predicting the response. The R value was 0.977, that was close to 1 indicating a great agreement between the experimental results and the theoretical values predicted by the model equation.

The optimization of the extracellular chitosanase production using RSM increased the enzyme production up to 13.9U/mL which was slightly higher than the predicted value (13.126U/

mL) (Table 3) which produced by the use of the medium composed of (g/L) chitosan, 30; K<sub>2</sub>HPO<sub>4</sub>, 1.5; MgSO<sub>4</sub>, 0.4; KCl, 4.0; yeast extract, 18.5 and FeSO<sub>4</sub>, 0.01; at pH 5.5, 30°C and 180rpm for 96h.

#### *Validation of the model*

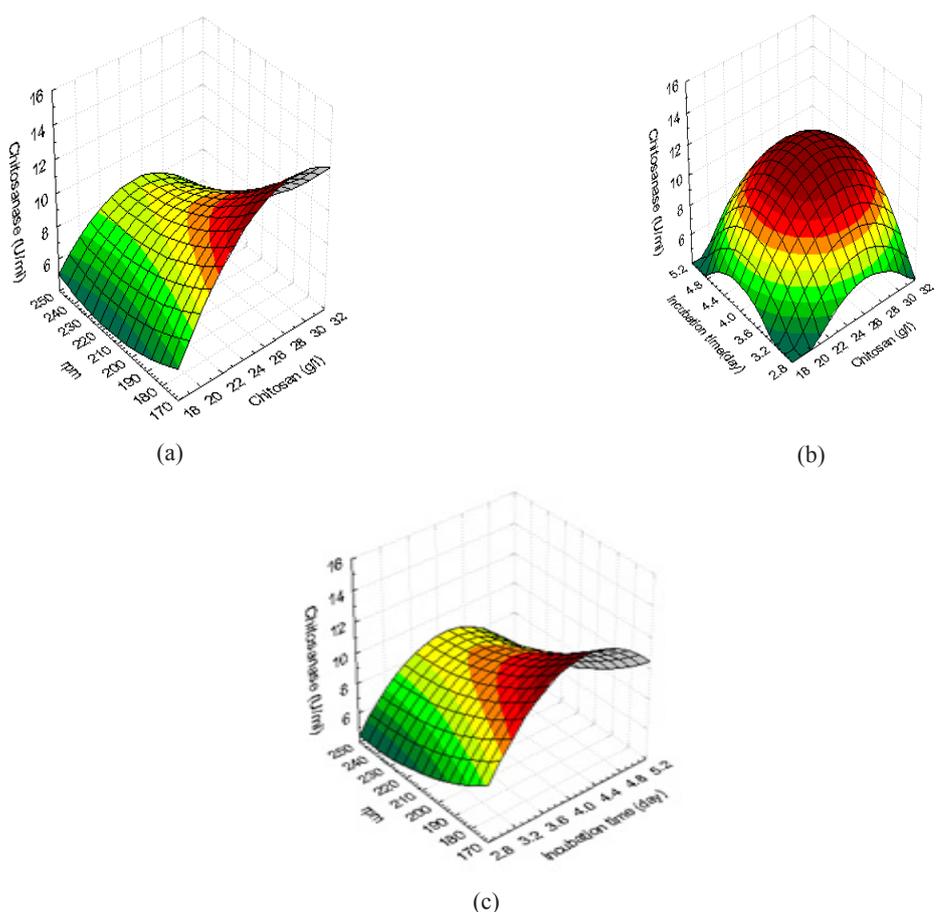
The validation was carried out under the optimum medium conditions in which the fermentation medium containing 3% chitosan concentration was used at 180rpm for 4days incubation period. The experimental chitosanase production of 13.9U/mL was obtained and it was close to its predicted value (13.126U/mL), validating the proposed model.

#### *Partial purification of chitosanase*

Partial purification of chitosanase by acetone and ammonium sulphate led to a low recovered chitosanase activity (unpublished data). The fraction at 60% ethanol had the highest recovered activity 47.16% with specific activity 16.7U/mg protein (4.91 times higher than the crude enzyme solution) but this fraction did not show a good stability so the fraction at 30-70% ethanol with specific activity 8.33U/mg protein had higher stability and was used in the next experiments.

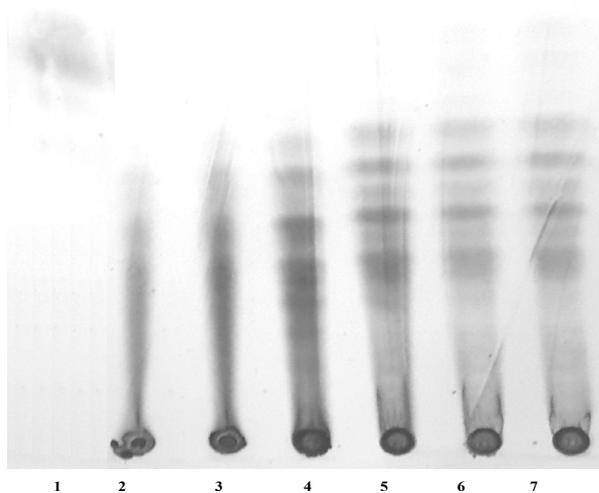
#### *Production of COS*

The TLC shown in Fig. 5 indicated that the highest hydrolysis yield of chitosan to COS (64.8%) was obtained with E/S 0.05U/mg at 55°C for 4h. These conditions were subjected



**Fig. 4. Response surface plot of chitosanase production**

- (a) Showing the interactive effects of different concentrations of chitosan (2-3%) and different rpm (180-240) at incubation time= 0.
- (b) Showing the interactive effects of different concentrations of chitosan (2-3%) and different incubation time (3-5days) at rpm= 0.
- (c) Showing the interactive effects of different rpm (180-240) and different incubation time (3-5days) at chitosan concentration= 0.



**Fig. 5. TLC plate of chitooligosaccharides produced by the hydrolysis of soluble chitosan by chitosanase. Lan 1: glucosamine. Lan 2-9: represent the hydrolysis products different E/S ratio 0.0125, 0.025, 0.05, 0.1, 0.15 and 0.2 respectively.**

*Box-Behnken design for the optimization of hydrolytic conditions*

Four variables (reaction time,  $X_1$ ; reaction temperature,  $X_2$ ; enzyme units,  $X_3$  and substrate concentration (%),  $X_4$ ) were chosen to determine their optimal response region in the hydrolysis of chitosan to chitoooligosaccharides. The amount of the released reducing sugar (RS) was used as an indicator for the hydrolysis of chitosan.

The variables were examined at three different levels, low (-), high (+) and control or basal (0) and the design matrix of the coded variables together with the experimental and the predicted results were shown in Table 5.

The regression analysis of the data (Table 6) resulted to a polynomial equation that was derived as follow:

$$Y_{RS(\%)} = 217.802 - 8.916 X_1 - 4.351 X_2 - 36.611 X_3 - 48.519 X_4 + 1.612 X_2^2 + 0.062 X_2^2 + 4.176 X_3^2 + 25.075 X_4^2 - 0.108 X_1 X_2 - 0.868 X_1 X_3 + 2.607 X_1 X_4 + 0.656 X_2 X_3 - 0.925 X_2 X_4 - 0.250 X_3 X_4 \quad \text{Eq. (8)}$$

where  $Y_{RS(\%)}$  is the response variable (amount of reducing sugar released %),  $X_1$  is the coded value of reaction time,  $X_2$  is the coded value of reaction temperature,  $X_3$  is the coded value of enzyme units and  $X_4$  is the coded value of substrate concentration.

**TABLE 5. Box- Behnken experimental design for the optimization of hydrolytic conditions.**

Trial	Independent variables				RS%	
	Time $X_1$ (h)	Temperature $X_2$ (°C)	Enzyme U $X_3$ (U)	Substrate conc. $X_4$ (%)	Observed	Predicted
1	-1 (3)	-1 (50)	0 (2)	0 (2)	55.6	55.417
2	-1 (3)	+1 (60)	0 (2)	0 (2)	74.1	71.210
3	+1 (5)	-1 (50)	0 (2)	0 (2)	74.1	64.203
4	+1 (5)	+1 (60)	0 (2)	0 (2)	90.7	82.157
5	0 (4)	0 (55)	-1 (1.5)	-1 (1.8)	61.1	58.481
6	0 (4)	0 (55)	-1 (1.5)	+1 (2.2)	74.1	63.907
7	0 (4)	0 (55)	+1 (2.5)	-1 (1.8)	70.4	69.837
8	0 (4)	0 (55)	+1 (2.5)	+1 (2.2)	83.3	75.163
9	-1 (3)	0 (55)	0 (2)	-1 (1.8)	64.8	61.127
10	-1 (3)	0 (55)	0 (2)	+1 (2.2)	66.7	64.418
11	+1 (5)	0 (55)	0 (2)	-1 (1.8)	68.5	72.037
12	+1 (5)	0 (55)	0 (2)	+1 (2.2)	70.4	73.242
13	0 (4)	-1 (50)	-1 (1.5)	0 (2)	48.1	56.558
14	0 (4)	-1 (50)	+1 (2.5)	0 (2)	64.8	56.558
15	0 (4)	+1 (60)	-1 (1.5)	0 (2)	66.7	66.912
16	0 (4)	+1 (60)	+1 (2.5)	0 (2)	81.5	81.495
17	-1 (3)	0 (55)	-1 (1.5)	0 (2)	48.1	56.292
18	-1 (3)	0 (55)	+1 (2.5)	0 (2)	68.5	69.335
19	+1 (5)	0 (55)	-1 (1.5)	0 (2)	61.1	65.291
20	+1 (5)	0 (55)	+1 (2.5)	0 (2)	72.2	80.070
21	0 (3)	-1 (50)	0 (2)	-1 (1.8)	55.6	56.918
22	0 (3)	-1 (50)	0 (2)	+1 (2.2)	55.6	64.144
23	0 (3)	+1 (60)	0 (2)	-1 (1.8)	70.4	72.400
24	0 (3)	+1 (60)	0 (2)	+1 (2.2)	66.7	75.926
25	0 (3)	0 (55)	0 (2)	0 (2)	64.8	64.800
26	0 (3)	0 (55)	0 (2)	0 (2)	64.8	64.800
27	0 (3)	0 (55)	0 (2)	0 (2)	64.8	64.800

TABLE 6. Analysis of Box-Behnken Design.

Term	Regression coefficient	Standard error	t- test	P-value
Intercept	217.802	859.732	0.253	0.804
$X_1$	-8.916	25.279	-0.353	0.730
$X_2$	-4.351	18.259	-0.238	0.816
$X_3$	-36.611	141.484	-0.259	0.800
$X_4$	-48.519	443.802	-0.109	0.915
$X_1^2$	1.612	1.019	1.582	0.140
$X_2^2$	0.062	0.152	0.406	0.692
$X_3^2$	4.176	15.220	0.274	0.788
$X_4^2$	25.075	91.992	0.273	0.790
$X_1X_2$	-0.108	0.332	-0.326	0.750
$X_1X_3$	-0.868	3.285	-0.264	0.796
$X_1X_4$	2.607	7.936	0.329	0.748
$X_2X_3$	0.656	2.053	0.319	0.755
$X_2X_4$	-0.925	4.200	-0.220	0.829
$X_3X_4$	-0.250	41.996	-0.006	0.995

F value = 2.007; P>F= 0.0097; R<sup>2</sup>=0.809; Adjusted R<sup>2</sup>=0.879

Three-dimensional response surfaces (Fig. 6a-f) were plotted on the basis of the model equation to investigate the interaction among the variables. Figure 7 showed the relation between the observed and the predicted results.

The optimization of the hydrolysis conditions increased the hydrolysis present from 64.8% to 90.7%. The analysis of the results using ANOVA showed a significant F-value (2.007) that indicated the significance of the model. Model terms had values of Prob> F (0.0097) less than 0.05, considered significant. The coefficient (R<sup>2</sup>) for the % of the released reducing sugar was 0.809 which indicated that the statistical model explained 80.9% of variability in the response.

#### Anticancer activity

The cytotoxic activities of the COS mixture eluted from TLC plate were tested on human liver cancer cell line (Hep-G2) and breast cancer

cell line (MCF7). The results showed in Fig. 8 indicated that the IC<sub>50</sub> of Hep-G2 was 12µg/mL while the IC<sub>50</sub> of MCF7 was 85.5µg/mL which means that the COS mixture had a stronger cytotoxic activity against Hep-G2 than MCF7.

#### Discussion

Chitosanases are specific enzymes used for the hydrolysis of chitosan through the hydrolysis of the β-(1-4)-glycosidic linkage to produce COS. The use of chitosanase to produce COS has several advantages. Many researchers are focused to find new microorganisms with high level of chitosanase production [1, 7]. In the current research the molecular identification of the chitosanase producer fungal isolate showed 99% similarity with *Dothideomyces* sp. Kirk *et al.*, [25] reported that *Dothideomyces* was the largest and the most diverse class of *Ascomycetes* and it was reported to produce chitosanase by

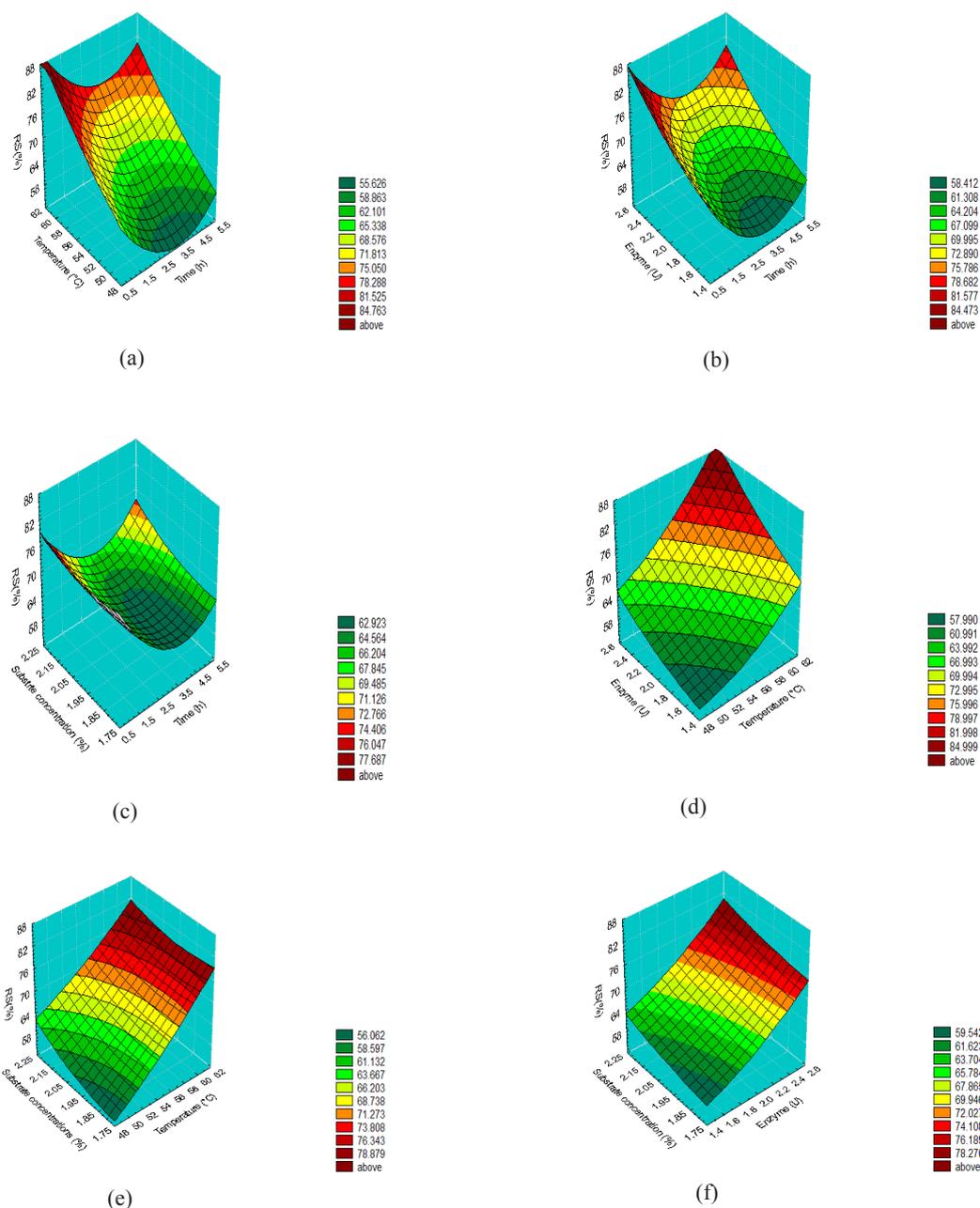


Fig. 6. Response surface plot of the amount of reducing sugar released %

- Showing the interactive effects of different time (3-5h) and different temperature (50-60°C) at  $X_3=0$  and  $X_4=0$ .
- Showing the interactive effects of different time (3-5h) and different enzyme units (1.5-2.5U) at  $X_2=0$  and  $X_4=0$ .
- Showing the interactive effects of different time (3-5h) and different substrate concentrations (1.8-2.2%) at  $X_2=0$  and  $X_3=0$ .
- Showing the interactive effects of different temperature (50-60°C) and different enzyme units (1.5-2.5U) at  $X_1=0$  and  $X_4=0$ .
- Showing the interactive effects of different temperature (50-60°C) and different substrate concentrations (1.8-2.2%) at  $X_1=0$  and  $X_3=0$ .
- Showing the interactive effects of different enzyme units (1.5-2.5U) and different substrate concentrations (1.8-2.2%) at  $X_1=0$  and  $X_2=0$ .

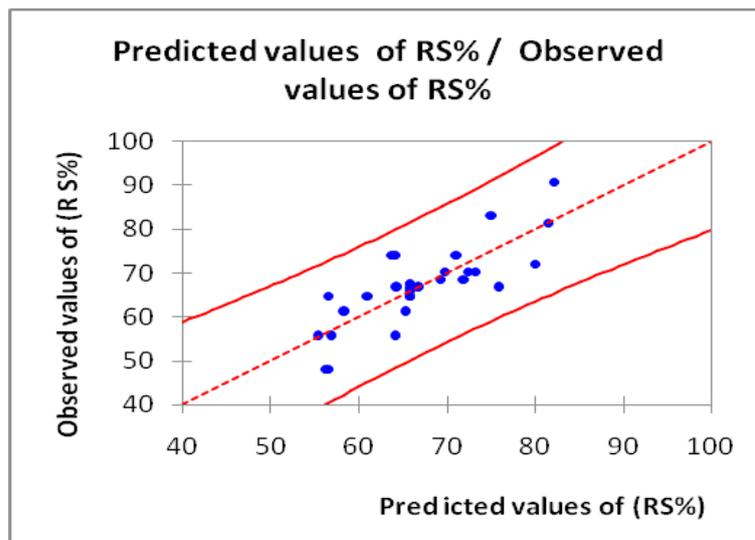


Fig. 7. Correlation between the observed and predicted values for RS (%) determined by the first-order polynomial equation.

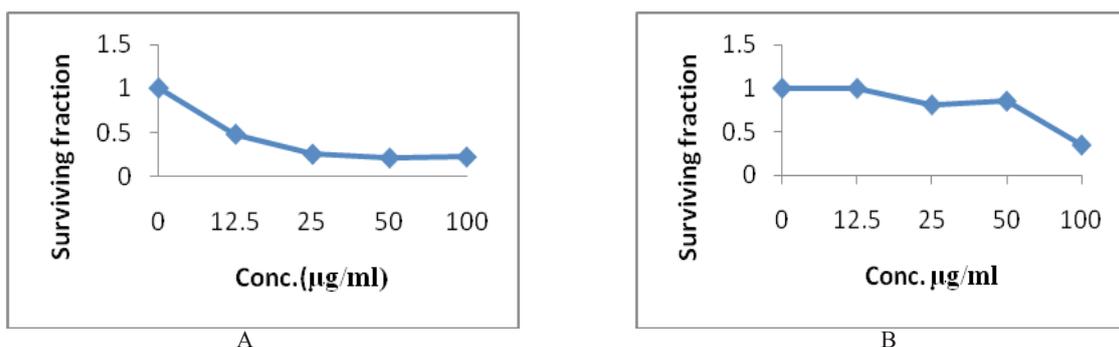


Fig. 8. Effect of different concentrations of COS on Hep-G2 (A) and MCF7 (B).

Ramos-Garza *et al.*, [26] and Zeiner *et al.*, [27].

The application of the statistical optimization of the microbial enzymes using RSM has received much attention [7-10]. So the optimization of the extracellular chitosanase production using RSM was carried out by two-phase multifactorial designs. The first phase involved the application of Plackett–Burman design which indicated that chitosan concentration, agitation speed and incubation period were the most significant variables. Thadathil & Velappan, [28] and Nidheesh *et al.*, [7] reported that chitosanase was an inducible enzyme and the incubation period was a strong variable that affects the production of the chitosanase from different microorganisms. The variables that showed the highest confidence level in the Plackett–Burman design were

further optimized by Box-Behnken design. RSM increased the enzyme activity up to 13.9U/mL which was 36.3% higher than that produced using the basal fermentation medium (10.2U/mL). The optimization of fungal chitosanases by using RSM has been reported by other authors. Zhang and Zhang, [29] reported the production of chitosanase with optimal activity 21.85U/mL using *Aspergillus fumigates* YT-1 by SmF while Nidheesh *et al.*, [7] reported the production of chitosanase with optimal activity 41.78U/g initial dry substrate using *Purpureocillium lilacinum* CFRNT12 by solid state fermentation.

The most interested application of chitosanase was the production of COS that recently known to be a desired molecules for various industries than chitosan due to their low molecular weight, low

viscosity and its water solubility [1, 2]. The most stable fraction of the partially purified enzyme was investigated for the hydrolysis of chitosan to COS and the statistical optimization of the hydrolysis conditions by applying Box-Behnken design increased the hydrolysis present up to 90.7%, which obtained after 5h with E/S 0.05U/mg at 60°C. Nidheesh *et al.*, [7] reported the use of chitosanase for the hydrolysis of chitosan with maximum COS production of 54.27%.

COS recently gained a great attention due to their various health benefits as antitumor, anti-inflammatory, anti-hypertensive and hypocholesterimic effects [3, 4]. The cytotoxic activities of the produced COS were tested on HEP-G2 and MCF7. The results indicated that the produced COS had a stronger cytotoxic activity against Hep-G2 (IC<sub>50</sub> was 12µg/mL) than MCF7 (IC<sub>50</sub> was 85.5µg/mL). The inhibitory effect of the COS on Hep-G2 was reported by Shen *et al.*, [30] and the anticancer activity of COS has been reported by other authors [4, 31, 32].

### Conclusion

In conclusion RSM using Plackett – Burman and Box-Behnken designs increased extracellular chitosanase production up to 13.9U/ml which was 36.3% higher than that produced using the basal fermentation medium (10.2U/mL). The statistical optimization of the hydrolysis of chitosan to COS with partially purified enzyme increased the hydrolysis products up to 90.7%. The produced COS had a stronger cytotoxic activity against Hep-G2 than MCF7.

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## تعظيم انتاج انزيم الكيتوزانيز بواسطة فطر الدسودوميستس وزيادة قدرته علي التكسير باستخدام الاسلوب الاحصائي

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تم استخدام الاسلوب الاحصائي لزيادة انتاج فطر الدسودوميستس لانزيم الكيتوزانيز. وجد ان تركيز الكيتوزان, سرعه الاهتزاز و مده التحضين هي اكثر العوامل المؤثره في انتاج انزيم الكيتوزانيز. كما استطاع الاسلوب الاحصائي لزيادة انتاج الانزيم لتصل الي 13.9 وحده/مل. وباستخدام الاسلوب الاحصائي لزيادة درجه تكسير الكيتوزان باستخدام الانزيم المنتج وجد ان افضل الظروف هي استخدام انزيم/ كيتوزان بنسبه 0.05 وحده/مل جرام عند 60 درجه حراره لمده 5 ساعات.